# An Introduction to the Fractional Brownian Motion Exposition & Insights

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UNIVERSITY OF TORONTO

#### CMS Seminar





1 Motivation of the Classical Process

2 Review of Gaussian Processes

The Fractional Brownian Motion & First Properties

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#### Motivation

Consider a finite collection of i.i.d. random variables  $X_1, ..., X_n$  with  $EX_i = 0$  and  $EX_i^2 = 1$ . Define a process  $(Z_t)_{t \in [0,1]}$  by  $Z\left(\frac{k}{n}\right) = X_k$  for  $k \in \{1, ..., n\}$  and linearly interpolate on intervals of the form  $\left[\frac{k}{n}, \frac{k+1}{n}\right]$ .

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Motivation of the Classical Process

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## **Definition:** Gaussian Process

A **Gaussian process** is a stochastic process  $(X_i)_{i \in T}$  such that any finite linear combination of the variables  $X_i$ ,  $t \in T$  is Gaussian.

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## Definition: Gaussian Process

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## Theorem: Characterization of Gaussian Processes

Gaussian processes are completely determined/characterized by their mean and covariance functions:

 $m: T \to \mathbb{R} \qquad , \qquad \Gamma: T \times T \to \mathbb{R}$  $t \mapsto E[X_{\cdot}] \qquad (s,t) \mapsto \operatorname{Cov}(X_{\cdot}, X_{t}).$ 

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# Proof:

For a Gaussian process  $(X_t)_{t \in T}$  with mean and covariance functions  $m_{\mathbf{X}}$ ,  $\Gamma_{\mathbf{X}}$  and any choice  $t_1 < t_2 < \cdots < t_p$  we have  $\mathbf{X} = (X_{t_1}, \dots, X_{t_p})$ ,  $t_i \in T$ , to be an  $\mathbb{R}^p$ -valued random variable with a multivariate normal distribution and hence to have characteristic function (Fourier transform of distribution):

$$\varphi_{\mathbf{X}}(u) = E[e^{i\langle u, (X_{t_1}, \dots, X_{t_p})\rangle}] = \exp(i\langle u, \mu_{\mathbf{X}} \rangle - \frac{1}{2}\langle u, \boldsymbol{\Sigma}_{\mathbf{X}} u \rangle),$$

where  $\mu = (E[X_{t_1}], \dots, E[X_{t_p}]) = (m(t_j))_{1 \le j \le p}$  is the mean vector and  $\Sigma_{i,j} = (\Gamma(t_i, t_j))_{1 \le i, j \le p}$  the covariance matrix.

#### <u>Proof:</u>

For a Gaussian process  $(X_t)_{t \in T}$  with mean and covariance functions  $m_X$ ,  $\Gamma_X$  and any choice  $t_1 < t_2 < \cdots < t_p$  we have  $\mathbf{X} = (X_{t_1}, \dots, X_{t_p}), t_i \in T$ , to be an  $\mathbb{R}^p$ -valued random variable with a multivariate normal distribution and hence to have characteristic function (Fourier transform of its distribution):

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where  $\mu_{\mathbf{X}} = (E[X_{t_1}], \dots, E[X_{t_p}]) = (m_{\mathbf{X}}(t_j))_{1 \le j \le p}$  is the mean vector and  $(\Sigma_{\mathbf{X}})_{i,j} = (\Gamma_{\mathbf{X}}(t_i, t_j))_{1 \le i, j \le p}$  the covariance matrix. Now, considering another Gaussian process  $(Y_t)_{t \in T}$  with the same mean and covariance functions as  $(X_t)_{t \in T}$  i.e.  $\Gamma_{\mathbf{Y}} = \Gamma_{\mathbf{X}}$ , and  $m_{\mathbf{X}} = m_{\mathbf{Y}}$ , we have the characteristic function of its finite dimensional marginal,  $(Y_{t_1}, \dots, Y_{t_p})$ , to be given identically as to that of above.

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#### <u>Proof:</u>

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Hence, we have an agreement in distribution of the finite dimensional marginals of  $(X_t)_{t \in T}$  and  $(Y_t)_{t \in T}$  and hence an equivalence in law.

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i.e. Gaussian processes sharing the same mean and cov. functions are equal in law.

# Moral Conclusion:

Specifying Gaussian processes amounts to specifying a mean and covariance structure on the collection. Similarly as to how specifying a Gaussian r.v. requires only a mean and variance!

# Definition: Centered Gaussian Process

A **centered Gaussian process** is a stochastic process  $(X_i)_{i \in T}$  such that any finite linear combination of the variables  $X_i$ ,  $t \in T$  is centered Gaussian.  $(t \mapsto E[X_i] \equiv 0)$ 

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#### Definition: Brownian motion

A stochastic process  $(B_t)_{t \ge 0}$  on  $(\Omega, \mathcal{F}, P)$  is a standard **Brownian Motion** if it satisfies one of the following three equivalent assertions:

(i)  $(B_t)_{t\geq 0}$  is a centered Gaussian process with covariance:  $Cov(B_s, B_t) = min(s, t) \coloneqq s \wedge t$ .

(ii)  $B_0 = 0$  a.s. and for every  $0 \le s < t$  the random variable  $B_t - B_s$  is independent of  $\sigma(B_r : r \le s)$  and  $B_t - B_s \sim \mathcal{N}(0, t - s)$ .

(iii)  $B_0 = 0$  a.s. and for  $0 = t_0 < t_1 < \dots < t_n$  the r.v.'s  $(B_{t_i} - B_{t_{i-1}})$  for  $1 \le i \le n$  are indep. and  $B_{t_i} - B_{t_{i-1}} \sim \mathcal{N}(0, t_i - t_{i-1})$ .

and as well:

 $(iv)(B_t)_{t\geq 0}$  has surely continuous sample paths:  $\forall \omega \in \Omega, t \mapsto B_t(\omega) \in C(\mathbb{R}_{\geq 0}, \mathbb{R}).$ 

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# Note:

Independence of increments are a key feature in the classical process!

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From here on out we consider only  $T = [0, \infty)$  i.e. positive time.

#### **Definition:** Fractional Brownian Motion

A **fractional Brownian Motion** of *Hurst* parameter/index  $H \in (0, 1]$  is a centered Gaussian process  $(B_t^H)_{t\geq 0}$  with covariance function

$$E[B_t^H B_s^H] = \frac{1}{2}(t^{2H} + s^{2H} - |t - s|^{2H}).$$

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#### Proposition:

If H = 1/2 then fBm is nothing but a classical Brownian motion.

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#### Proposition:

If H = 1/2 then fBm is nothing but a classical Brownian motion.

#### Proof:

Fix  $0 \le s < t$  and observe

$$E[B_s^{1/2}B_t^{1/2}] = \frac{1}{2}(t+s-|t-s|) = s = s \wedge t.$$

# Properties

# Self-Similarity

For  $\epsilon > 0$  given and  $H \in (0, 1)$ , we have  $(\epsilon^{-H} B_{\epsilon t}^{H})_{t \ge 0}$  is an fBm of Hurst index H.

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Illustration for H = 1/2:



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# Proof:

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# Proof: Exercise.

For  $\varepsilon > 0$  given and  $H \in (0, 1)$ , the process  $(\varepsilon^{-H} B^H_{\epsilon t})_{t \ge 0}$  is an fBm of Hurst index H.

#### Stationary Increments

For  $s \ge 0$  fixed and  $H \in (0, 1)$ , the process  $(B_{s+t}^H - B_s^H)_{t\ge 0}$  is an fBm of Hurst index H.

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## Continuity of Sample Paths

An fBm  $(B_t^H)_{t\geq 0}$  admits a continuous modification. That is we have some process  $(X_t)_{t\geq 0}$  such that  $t \mapsto X_t \in C[0, \infty)$  (surely) and for all  $t \geq 0$ ,  $P(B_t^H = X_t) = 1$ .

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# Theorem: Kolmogorov-Čhenstov Continuity Theorem

Assume that for a stochastic process  $(X_t)_{t \ge 0}$  there exists K > 0, p > 0,  $\beta > 0$  such that for all  $s, t \ge 0$ :

$$E[|X_t - X_s|^p] \le K|t - s|^{1+\beta}.$$

Then the process has a continuous modification, i.e. a process  $(\widetilde{X}_t)_{t\geq 0}$  such that  $t \mapsto \widetilde{X}_t \in C[0, \infty)$  and for all  $t \geq 0$   $P(X_t = \widetilde{X}_t) = 1$ .

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#### Proof:

Simply observe:

$$E[(B_t^H - B_s^H)^2] = |t - s|^{2H},$$

and apply Kolmogorov-Čhenstov.

For  $\varepsilon > 0$  given and  $H \in (0, 1)$ , the process  $(\varepsilon^{-H} B^H_{\varepsilon t})_{t \ge 0}$  is an fBm of Hurst index H.

#### Stationary Increments

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## The fBm is (in general) not Markov

Let  $(B_t^H)_{t\geq 0}$  be a fractional Brownian motion of Hurst index  $H \neq 1/2$ . Then  $(B_t^H)_{t\geq 0}$  is not a Markov process.

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Then how do we describe the dependence structure of fBm and how does such structure vary with the Hurst index?

Disjoint increments of an fBm of Hurst index  $H \in (0, 1]$  are *negatively* correlated for  $H \in (0, 1/2)$  and *positively* correlated for  $H \in (1/2, 1)$ .

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Suppose that  $0 \le s_1 < t_1 < s_2 < t_2$  so as to ensure  $[s_1, t_1] \cap [s_2, t_2] = \emptyset$ . Then one can check the covariance of the increments to be given as:

$$E[(B_{t_1}^H - B_{s_1}^H)(B_{t_2}^H - B_{s_2}^H)] = \frac{1}{2} \Big( |t_2 - s_1|^{2H} - |t_2 - t_1|^{2H} - \Big( |s_2 - s_1|^{2H} - |s_2 - t_2|^{2H} \Big) \Big).$$

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Suppose that  $0 \le s_1 < t_1 < s_2 < t_2$  so as to ensure  $[s_1, t_1] \cap [s_2, t_2] = \emptyset$ . Then one can check the covariance of the increments to be given as:

$$E[(B_{t_1}^H - B_{s_1}^H)(B_{t_2}^H - B_{s_2}^H)] = \frac{1}{2} \Big( |t_2 - s_1|^{2H} - |t_2 - t_1|^{2H} - \Big( |s_2 - s_1|^{2H} - |s_2 - t_2|^{2H} \Big) \Big).$$

Then considering the map  $f(x) = x^{2H}$  and putting  $a_1 = t_2 - s_1, a_2 = t_2 - t_1, b_1 = s_2 - s_1, b_2 = s_2 - t_1$  gives that  $a_1 - a_2 = b_1 - b_2 = t_1 - s_1$  (note that  $b_2 < a_2 < b_1 < a_1$ ) and allows the above to be expressed as follows:

$$E[(B_{t_1}^H - B_{s_1}^H)(B_{t_2}^H - B_{s_2}^H)] = \frac{1}{2}(f(a_1) - f(a_2) - (f(b_1) - f(b_2))).$$

Disjoint increments of an fBm of Hurst index  $H \in (0, 1]$  are *negatively* correlated for  $H \in (0, 1/2)$  and *positively* correlated for  $H \in (1/2, 1)$ .

#### Proof:

Suppose that  $0 \le s_1 < t_1 < s_2 < t_2$  so as to ensure  $[s_1, t_1] \cap [s_2, t_2] = \emptyset$ . Then one can check the covariance of the increments to be given as:

$$E[(B_{t_1}^H - B_{s_1}^H)(B_{t_2}^H - B_{s_2}^H)] = \frac{1}{2} \Big( |t_2 - s_1|^{2H} - |t_2 - t_1|^{2H} - \Big( |s_2 - s_1|^{2H} - |s_2 - t_2|^{2H} \Big) \Big).$$

Then considering the map  $f(x) = x^{2H}$  and putting  $a_1 = t_2 - s_1, a_2 = t_2 - t_1, b_1 = s_2 - s_1, b_2 = s_2 - t_1$  gives that  $a_1 - a_2 = b_1 - b_2 = t_1 - s_1$  (note that  $b_2 < a_2 < b_1 < a_1$ ) and allows the above to be expressed as follows:

$$E[(B_{t_1}^H - B_{s_1}^H)(B_{t_2}^H - B_{s_2}^H)] = \frac{1}{2}(f(a_1) - f(a_2) - (f(b_1) - f(b_2))).$$

Now as f'' < 0 for  $H \in (0, 1/2)$  we have for such H that  $E[(B_{t_1}^H - B_{s_1}^H)(B_{t_2}^H - B_{s_2}^H)] < 0$ . And since f'' > 0 for  $H \in (1/2, 1)$  we have for such H that  $E[(B_{t_1}^H - B_{s_1}^H)(B_{t_2}^H - B_{s_2}^H)] > 0$ .

# Sample Path Regularity w.r.t H



FIGURE 1. Paths of fBm for different values of H.

## Gaussian Beam; H = 1

One can check that for H = 1 we have  $(B_t^H)_{t \ge 0} \stackrel{d}{=} (t\xi)_{t \ge 0}$ , for  $\xi \sim \mathcal{N}(0, 1)$ .

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# Limiting Case

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